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INFINITE TOPOLOGICALLY RANDOM CHANNEL NETWORKS¹

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ABSTRACT

Individual channel networks ordinarily are portions of far larger, essentially infinite, networks. The all network is by definition at infinite topologically random network if the populations of subnetworks within it are topologically random. From such a network, the probability of randomly drawing a link, a subnetwork or a basin with Strahler order ω is $1/2^\omega$; and that of randomly drawing a stream of order ω is $3/4^\omega$. The probability of drawing a link of magnitude μ , that is, one having μ sources ultimately tributary to it, is equal to the probability of a first passage through the origin at step $2\mu - 1$ in a symmetric random walk, a fact which suggests a useful mathematical analogy between random walks and infinite topologically random networks. Assuming uniform link length equal to the constant of channel maintenance, which in turn is the reciprocal of drainage density, the probability distributions for links and streams of various orders may be interpreted as crude geomorphological "laws" analogous to Horton's laws of drainage composition. These distributions predict geometric-series "laws" in which, using Strahler orders, the bifurcation ratio is $1/2$, the link-number ratio is $1/2$, the length ratio is 2, the cumulative-length ratio is 4, and the basin-area ratio is 4, all in agreement with the observed ratios. They also predict values of $4/3$ and $3/2$, respectively, for the dimension ratios of total number of Strahler streams to network magnitude and of Strahler stream frequency to square of the drainage density, in agreement with the values of 1.34 and 0.694 found empirically.

INTRODUCTION

Individual channel networks in nature do not ordinarily exist independently, but are portions of far larger networks that for practical purposes are essentially infinite. The hypothesis that in the absence of geologic controls natural networks will be topologically random (Shreve, 1966, p. 27) thus implicitly contains the corollary that in the over-all infinite network, which will be termed an *infinite topologically random channel network*, all topologically distinct subnetworks with the same number of sources occur with equal frequency. The purposes of this paper are to deduce some of the geomorphologically significant probability distributions associated with these infinite topologically random networks, to demonstrate the close connection of such networks with random walks, and to derive theoretically several of the well-known empirical "laws" of geomorphology.

DEFINITIONS

Most of the specialized terms used in this paper—such as *link*, *fork*, *source*, *topologi-*

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cally identical, topologically distinct, topologically random—have been defined in a previous publication (Shreve, 1966, p. 27). Strahler stream orders (1952, p. 1120; Shreve, 1966, p. 21–22) will be used exclusively. The terms *channel network* or, for convenience, simply *network* will refer to the idealized concept of channel network (Shreve, 1966, p. 27) in which one, and only one, path exists between any two points in which at its upstream end each link either connects to two other links or terminates at a source. Stated negatively, and somewhat less precisely, no lakes with multiple inlets, no islands of major extent, and no confluences of more than two channels at a single place are permissible in the idealized channel networks considered in this paper. The terms *exterior link* and *interior link* will signify links terminating at their upstream ends in sources and forks, respectively; these terms, Melton's relationships (1952, p. 345; Shreve, 1966, p. 27) state that every idealized network with μ sources will have $2\mu - 1$ links, of which μ are exterior links and $\mu - 1$ are interior links.

The phrase *link drawn at random* will signify a link selected in such a way that all links in the specified target population are equally likely to be drawn, whether which may be either real or hypothetical.

CHANNEL NETWORKS¹

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 Graybill, 1963, p. 141–142). It is com-
 pletely equivalent to the phrases *network
 drawn at random* or *basin drawn at random*,
 because each link of given magnitude or
 order or both in a population of networks is
 the outlet of a single subnetwork of the same
 magnitude or order or both. Actually draw-
 ing links or networks at random from natu-
 ral populations is not easily accomplished;
 nevertheless, for meaningful statistical anal-
 ysis it is an absolute necessity.

CONCEPT OF MAGNITUDE

The *magnitude* of a link in a channel net-
 work is herein defined by the following two
 rules. (1) Each exterior link has magnitude
 1. (2) If links of magnitude μ_1 and μ_2 join,
 then the resultant link downstream has
 magnitude $\mu_1 + \mu_2$. The application of these
 rules to a typical channel network is illus-
 trated in figure 1. By induction, the mag-
 nitude of a link is equal to the total number of
 sources ultimately tributary to it. Magni-
 tude is thus a purely topological concept, for
 it involves only the interconnections and not
 the lengths, shapes, or orientations of the
 links comprising the network. Like the sys-
 tems of stream ordering of Horton (1945, p.
 281), Strahler (1952, p. 1120 n.), and, more
 recently, Scheidegger (1965, p. B188), the
 concept of magnitude assumes the existence
 of sources, that is, of objectively describable
 points that in fact or by definition are the
 points farthest upstream in a channel net-
 work.

In analogy with basin order, the mag-
 nitude of a channel network or drainage basin
 is equal to the magnitude of its highest mag-
 nitude link. Thus, if two networks are tribu-
 tary to the same link, the resultant network
 has magnitude equal to the sum of the mag-
 nitudes of the two tributary networks. Net-
 works with equal magnitudes have equal
 numbers of links, forks, sources, Horton
 streams, and first-order Strahler streams,
 and are therefore comparable in topological
 complexity (Shreve, 1966, p. 27). A reason-
 able conjecture is that they are comparable
 in other ways as well.

Despite its apparent complexity, the sys-
 tem of stream ordering proposed by
 Scheidegger (1965) is simply related to the
 concept of magnitude. If the Scheidegger
 order and the magnitude of a link (or net-
 work) are denoted by X and μ , respectively,
 then

$$X = \log_2 2^\mu, \quad (1)$$

in which \log_2 signifies the logarithm to base
 2.

According to Scheidegger (1965, p. B188–
 B189), if the second-order streams as classi-
 fied by his system “are treated as first-order

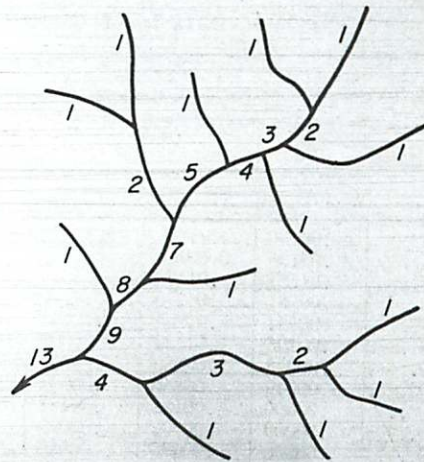


FIG. 1.—Link magnitudes for a typical channel network.

streams, then all orders are simply reduced
 by one.” Equation (1) makes it easy to see,
 however, that this statement is incorrect,
 unless the previous first-order streams are
 treated as zero-order streams that contrib-
 ute to the order of the higher-order streams
 even though, as Scheidegger himself states
 (1965, p. B188), they are in fact non-existent
 in nature. Thus, one of the main advantages
 of the Horton and Strahler systems (Shreve,
 1966, p. 22–23) is lost in the Scheidegger
 system.

PROBABILITY DISTRIBUTIONS

The probability $p(\mu, \omega)$ that a link
 drawn at random from an infinite topologi-

cally random network will have specified magnitude μ and Strahler order ω can be computed from the recursive relationship

$$p(\mu, \omega) = \frac{1}{2} \sum_{\alpha=1}^{\mu-1} [p(\alpha, \omega-1) \times p(\mu-\alpha, \omega-1) + 2p(\alpha, \omega) \sum_{\beta=1}^{\omega-1} p(\mu-\alpha, \beta)], \quad (2)$$

$$p(1, 1) = \frac{1}{2}, \quad p(\mu, 1) = 0,$$

$$p(1, \omega) = 0, \quad \mu, \omega = 2, 3, \dots$$

can be derived either in the same direct fashion or by summation using (2). Proceeding directly,

$$u(\omega) = \frac{1}{2} \left[u(\omega-1)^2 + 2u(\omega) \sum_{\beta=1}^{\omega-1} u(\beta) \right], \quad (3)$$

$$u(1) = \frac{1}{2}, \quad \omega = 2, 3, \dots,$$

$$v(\mu) = \frac{1}{2} \sum_{\alpha=1}^{\mu-1} v(\alpha) v(\mu-\alpha), \quad (4)$$

$$v(1) = \frac{1}{2}, \quad \mu = 2, 3, \dots$$

TABLE 1
INFINITE TOPOLOGICALLY RANDOM CHANNEL NETWORKS

μ	$p(\mu, \omega)$							$v(\mu)$
	$\omega=2$	$\omega=3$	$\omega=4$	$\omega=5$	$\omega=6$	$\omega=7$	$\omega=8$	
1.....								0.50000
2.....	0.12500							.12500
3.....	.06250							.06250
4.....	.03125	0.00781						.03906
5.....	.01562	.01172						.02734
6.....	.00781	.01270						.02051
7.....	.00391	.01221						.01611
8.....	.00195	.01111	0.00003					.01309
9.....	.00098	.00983	.00011					.01091
10.....	.00049	.00856	.00022					.00927
20.....	.00000	.00181	.00141	0.00000				.00322
50.....	.00000	.00002	.00070	.00009	0.00000			.00080
100.....	.00000	.00000	.00010	.00018	.00000	0.00000		.00023
200.....	.00000	.00000	.00000	.00009	.00001	.00000	0.00000	.00010
500.....	.00000	.00000	.00000	.00000	.00002	.00000	.00000	.00002
1000.....	0.00000	0.00000	0.00000	0.00000	0.00001	0.00000	0.00000	0.00001
$u(\omega)$	0.25000	0.12500	0.06250	0.03125	0.01562	0.00781	0.00391

The factor $\frac{1}{2}$ is the probability of drawing an interior link; and the sum on α accounts for the various ways that pairs of smaller networks can be tributary to this link to give it the specified μ and ω . Recursive relationships for the marginal probabilities

$$u(\omega) = \sum_{\mu=1}^{\infty} p(\mu, \omega)$$

and

$$v(\mu) = \sum_{\omega=1}^{\infty} p(\mu, \omega)$$

As before, the factor $\frac{1}{2}$ is the probability of drawing an interior link; and the remaining factor accounts for the various ways the pairs of smaller networks can be tributary to this link to give it the specified μ or ω .

Representative values of $p(\mu, \omega)$, $u(\omega)$ and $v(\mu)$ are presented in table 1. For any particular magnitude the distribution of probabilities with order has a relatively sharp peak, whereas for any particular order the distribution with magnitude has a relatively broad peak; hence, in a topologically random population magnitude more pre

ived either in the same direct
y summation using (2). Proceed

$$v(\omega-1)^2 + 2u(\omega) \sum_{\beta=1}^{\omega-1} u(\beta) \Big], \quad (3)$$

$$v(1) = \frac{1}{2}, \quad \omega = 2, 3, \dots,$$

$$v(\mu) = \frac{1}{2} \sum_{\alpha=1}^{\mu-1} v(\alpha) v(\mu - \alpha), \quad (4)$$

$$v(\mu) = \frac{1}{2}, \quad \mu = 2, 3, \dots$$

NEL NETWORKS

		$v(\mu)$
	$\omega=7$	$\omega=8$
00		0.50000
01		.12500
02		.06250
03		.03906
04		.02734
05		.02051
06		.01611
07		.01309
08		.01091
09		.00927
10		.00822
11		.00780
12		.00728
13	0.00000	0.00000
14	.00000	.00010
15	.00000	.00002
16	0.00000	0.00001
62	0.00781	0.00391

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n interior link; and the remaining
counts for the various ways that
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ad peak; hence, in a topologically
population magnitude more pre-

cisely characterizes a network than does
order.

A closed expression for $u(\omega)$ can readily
be derived from (3). Summation of the first
 $\omega - 1$ terms of $u(\omega)$ gives, using (3),

$$2 \sum_{\beta=1}^{\omega-1} u(\beta) = \left[\sum_{\beta=1}^{\omega-1} u(\beta) \right]^2 - u(\omega - 1)^2 + 1. \quad (5a)$$

Completing the square, taking the square
root, and substituting into (3) leads to the
relationship

$$2u(\omega) = u(\omega - 1), \quad (5b)$$

$$u(1) = \frac{1}{2}, \quad \omega = 2, 3, \dots,$$

from which

$$u(\omega) = 1/2^\omega, \quad \omega = 1, 2, \dots, \quad (5c)$$

follows by mathematical induction.

A closed expression for $v(\mu)$ can be de-
rived from (4) by means of the generating
function $R(r)$, where

$$R(r) = \sum_{\mu=1}^{\infty} v(\mu) r^\mu. \quad (6a)$$

Squaring both sides of (6a) and collecting
the coefficients of like powers of r gives

$$R^2 = \sum_{\mu=2}^{\infty} r^\mu \sum_{\alpha=1}^{\mu-1} v(\alpha) v(\mu - \alpha); \quad (6b)$$

hence, using (4),

$$R^2 = 2R - r, \quad (6c)$$

$$R = 1 - (1 - r)^{1/2},$$

from which

$$v(\mu) = \frac{2^{-(2\mu-1)}}{2\mu-1} \binom{2\mu-1}{\mu}, \quad (6d)$$

$$\mu = 1, 2, \dots,$$

follows by the binomial expansion. The
minus sign is chosen in the quadratic formu-
la because the power-series expansion in
(6a) represents the branch of the curve of R
versus r that passes through the origin.

The distribution $v(\mu)$ can also be derived

from the probability $w(\mu; M)$ of drawing a
link of magnitude μ at random from a topo-
logically random population of networks of
magnitude M . This probability is

$$w(\mu; M) = \frac{(M-\mu+1)N(M-\mu+1)N(\mu)}{(2M-1)N(M)}, \quad (7a)$$

where $N(\mu)$ is the number of topologically
distinct networks of magnitude μ . The de-
nominator is equal to the total number of
links of all magnitudes in the population of
 $N(M)$ networks of magnitude M ; and the
numerator is therefore the number of links
of magnitude μ in this population, which in
turn is the number of ways that the $N(\mu)$
networks of magnitude μ can be attached to
the $M - \mu + 1$ sources of each of the
 $N(M - \mu + 1)$ networks of magnitude
 $M - \mu + 1$ to form networks of magnitude
 M .

Substitution of Cayley's closed expression
for $N(\mu)$,

$$N(\mu) = \frac{1}{2\mu-1} \binom{2\mu-1}{\mu} \quad (7b)$$

(Shreve, 1966, p. 29), gives

$$w(\mu; M) = \frac{1}{2\mu-1} \binom{2\mu}{\mu} \binom{2(M-\mu)}{M-\mu} / \binom{2M}{M}, \quad (7c)$$

from which (6d) follows by taking the limit
as M approaches infinity.

Representative values of $w(\mu; M)$ are
presented in table 2.

For any network magnitude M the great-
est (or modal) probability is that of drawing
a link of magnitude 1, inasmuch as

$$w(1; M) = M/(2M - 1) \geq \frac{1}{2}. \quad (8)$$

At the other end of the distribution, how-
ever, for $M > 3$ the probability of drawing
a link of magnitude M is greater than that of
drawing one of magnitude $M - 1$. This is
because networks of magnitude M always
have one link of that magnitude, whereas
they very often have no links of the next
smaller magnitude and very rarely have

two, which is the maximum possible. Setting the ratio of two successive terms equal to 1 and solving for μ shows that the minimum in the distribution is located approximately at $\mu = 3(M + 1)/4$.

The probability $s(\omega)$ of drawing a Strahler stream of order ω at random from the streams (not links) comprising an infinite topologically random network can be derived from (5c) and (6d) by noting that, because in any network the links whose tributaries are both of order $\omega - 1$ are in one-to-

could readily be tested against natural populations of channel networks by means of the goodness-of-fit test (see, e.g., Siegel, 1956, p. 42-52, 59-60, and Mood and Graybill, 1963 p. 308-309).

CONNECTION WITH RANDOM WALKS

A symbolic representation of any channel network can be constructed as follows. Start at the outlet and traverse the network always turning left at forks and reversing direction at sources, until the outlet is again

TABLE 2
POPULATIONS OF FINITE TOPOLOGICALLY RANDOM CHANNEL NETWORKS

M	$w(\mu;M)$							
	$\mu=1$	$\mu=2$	$\mu=3$	$\mu=4$	$\mu=5$	$\mu=6$	$\mu=7$	$\mu=8$
1.....	1.00000							
2.....	0.66667	0.33333						
3.....	0.60000	.20000	0.20000					
4.....	0.57143	.17143	.11429	0.14286				
5.....	0.55556	.15873	.09524	.07936	0.11111			
6.....	0.54545	.15152	.08658	.06494	.06061	0.09091		
7.....	0.53846	.14685	.08158	.05828	.04895	.04895	0.07692	
8.....	0.53333	.14359	.07832	.05439	.04351	.03916	.04103	0.06667
9.....	0.52941	.14118	.07602	.05183	.04031	.03455	.03258	.03529
10.....	0.52632	.13932	.07430	.05001	.03819	.03183	.02858	.02786
20.....	0.51282	.13167	.06772	.04360	.03151	.02445	.01992	.01683
50.....	0.50505	.12756	.06445	.04072	.02882	.02185	.01737	.01428
100.....	0.50251	.12627	.06346	.03987	.02805	.02115	.01671	.01365
200.....	0.50125	.12563	.06297	.03946	.02769	.02082	.01640	.01336
500.....	0.50050	.12525	.06269	.03922	.02748	.02063	.01623	.01320
1000.....	0.50025	.12512	.06259	.03914	.02741	.02057	.01617	.01314
∞	0.50000	0.12500	0.06250	0.03906	0.02734	0.02051	0.01611	0.01309

one correspondence with the streams of order ω ,

$$s(2)/s(1) = v(2)/v(1) = \frac{1}{4} \quad (9a)$$

and

$$s(\omega + 1)/s(\omega) = \frac{1}{2}u(\omega)^2 / \frac{1}{2}u(\omega - 1)^2 = \frac{1}{4}, \quad \omega = 2, 3, \dots \quad (9b)$$

These equations define a geometric series with ratio $\frac{1}{4}$, from which

$$s(\omega) = 3/4^\omega, \quad \omega = 1, 2, \dots, \quad (9c)$$

follows by mathematical induction.

The distributions (5c), (6d), and (9c)

reached. During the traverse, generate a sequence of I 's and E 's by recording an I the first time a given interior link is traversed and an E the first time a given exterior link is traversed. Each link will be traversed twice but recorded only once.

If a right turn instead of a left turn is made at each fork, a different sequence will result, which is the sequence for the mirror-image network. The new sequence will not be the reverse of the original, however; and, in general, symmetry in the network will not be evident in the sequence, and vice versa. The reverse sequence is generated by turn-

ily be tested against natural popu-
channel networks by means of the
f-fit test (see, e.g., Siegel, 1956, p.
60, and Mood and Graybill, 1963,
).

SECTION WITH RANDOM WALKS

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can be constructed as follows.
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ning left at forks and reversing
t sources, until the outlet is again

IM CHANNEL NETWORKS

	$\mu = 6$	$\mu = 7$	$\mu = 8$
11			
15	0.09091		
19		0.07692	
23	.04895	.04103	0.06667
27	.03916	.03258	.03529
31	.03455	.02858	.02786
35	.03183	.01992	.01683
39	.02445	.01737	.01428
43	.02185	.01671	.01365
47	.02115	.01640	.01336
51	.02082	.01623	.01320
55	.02063	.01617	.01314
59	.02057	.01611	.01309
63	0.02051		

uring the traverse, generate a se-
I's and *E*'s by recording an *I* the
a given interior link is traversed
he first time a given exterior link
d. Each link will be traversed
recorded only once.

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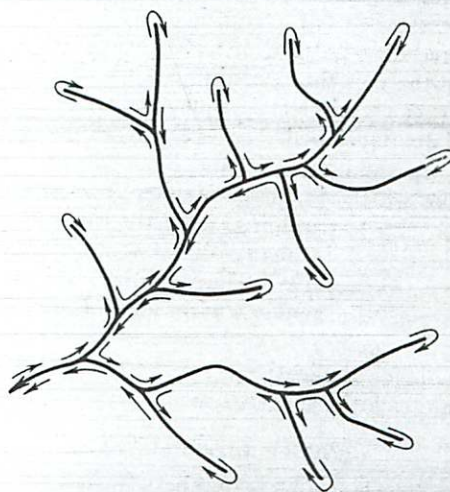
ing right and recording the *I*'s and *E*'s the
second time given links are traversed rather
than the first time. The mirror-image and
reverse sequences, although necessary for
certain types of investigation, will not be
used in this paper.

Topologically identical networks will
have identical sequences, and topologically
distinct networks will have different unique
sequences. Obviously, some possible se-
quences, such as all *I*'s or all *E*'s, cannot
correspond to channel networks. Sequences
corresponding to networks of more than one
link, for instance, always begin with *I* and
end with two successive *E*'s. More generally,
because of the fact that in any network or
subnetwork the number of interior links is
always exactly one less than the number of
exterior links, only those sequences are possi-
ble in which, as they are recorded, the
number of *E*'s never exceeds the number of
I's except at the terminal *E*. Put another
way, if the *I*'s count $+1$ and the *E*'s count
 -1 , then the partial sums can never be
negative, except for the last, which will be
 -1 . Thus, on a graph the curve of partial
sums may fall to the level of the origin, but
it does not drop below it until the terminal
E.

The steps from network to sequence to
graph are illustrated in figure 2.

Graphs like that of figure 2 occur widely
in the theory of random walks; and it is
worthwhile to compare their properties with
those of topologically random networks.
The graphs dropping below the axis for the
first time at step $2\mu - 1$ correspond to the
topologically distinct networks with $2\mu - 1$
links; hence, the number of such graphs is
 $N(\mu)$, as shown directly by Feller (1957, p.
71; see Shreve, 1966, p. 29, for derivation in
terms of networks). In random-walk ter-
minology $N(\mu)$ different one-dimensional
paths make a first passage of the origin at
step $2\mu - 1$. Following the analogy further,
if the graphs are generated by unbiased coin
tossing, counting heads as *I* and tails as *E*,
then the probability that the graph will drop
below the axis, that is, that a first passage

will occur, at step $2\mu - 1$ is equal to $v(\mu)$
(Feller, 1957, p. 73-75). Thus, the probabili-
ty that a link drawn at random from an
infinite topologically random network will
have magnitude μ , thereby defining a sub-
network with $2\mu - 1$ links, is exactly the
same as the probability that in a symmetric
random walk the first passage of the origin
will occur at step $2\mu - 1$ (Feller, 1957, p.
76).



I I E I I I E E I E I I I E E E E E I I I E E E E

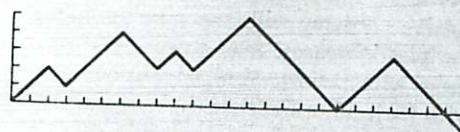


FIG. 2.—Symbolic and graphical representations
of a typical channel network.

A symmetric random walk can be re-
garded as the outcome of a sequence of
Bernoulli trials with probability of "suc-
cess" equal to $\frac{1}{2}$ (Feller, 1957, p. 135, 311);
hence, in the infinite sequence of *I*'s and
E's corresponding to an infinite topologi-
cally random network, *E* appears at any
specified position with probability $\frac{1}{2}$ regard-
less of the pattern anywhere else in the se-
quence. All possible subsequences are
equally likely to appear, corresponding to

the definition of topologically random networks (for proof of the zero probability of occurrence of "pathological" sequences, such as all E 's, see Feller, 1957, p. 189-197).

In the sequence corresponding to an infinite network that is not topologically random, on the other hand, although the two letters appear in equal numbers as required by Melton's relationships, the probability that E appears at a specified position is not necessarily $\frac{1}{2}$. Instead, it depends upon the preceding pattern of letters, corresponding to the fact that certain network topologies, hence certain letter sequences, are preferentially developed.

As random walks and Bernoulli trials occur widely in fields ranging from nuclear physics to gambling, their theory is highly developed in many different guises (see, e.g., Feller, 1957, chaps. iii, xi, xiii, xiv); hence, the analogy with infinite topologically random channel networks makes available for application a vast body of theorems and solutions.

GEOMORPHOLOGICAL "LAWS"

The law of large numbers (Feller, 1957, p. 141-142, 189-191) in conjunction with (9c) implies that, in random samples from an infinite topologically random network, the average number of Strahler streams of successive orders will, as the number of samples increases, tend toward a geometric series with ratio $\frac{1}{4}$. Similarly, in conjunction with (5c), it implies that the average number of links of successive orders will tend toward a geometric series with ratio $\frac{1}{2}$. In analogy with Horton's laws of drainage composition, these relationships might be termed the "law of stream numbers" and the "law of link numbers," respectively, although, like Horton's laws, they are not true statistical laws in the sense proposed in my paper on the law of stream numbers (Shreve, 1966, p. 17). Combining these two "laws" leads to the further relationship that the average number of links per stream increases with order as a geometric series with ratio 2.

Derivation of a "law of stream lengths"

requires a hypothesis concerning the lengths of individual links. The simplest, and one that does not appear to be too far from the truth, is that all of the links have the same length. In this case, the "law of stream lengths" corresponding to the previous crude "laws" would state that the average length of streams increases with order as a geometric series with ratio 2. In like manner the average total length of streams in basins of given order would increase with successive orders approximately as a geometric series with ratio 4, in good agreement with observation (e.g., Schumm, 1956, p. 604-605).

Suggestive as these geomorphological "laws" are, they are based upon averages over an infinite population and so, like Horton's laws (Shreve, 1966, p. 17), do not supply the complete distributions needed for statistical analysis of observations made on natural populations. Moreover, the hypothesis of constant link length amounts to using the average link length. A better hypothesis, for example, might be that the link lengths have a log-normal distribution (M. A. Melton, unpublished analysis; Schumm 1956, p. 607-608) with mean inversely proportional to drainage density, or perhaps with exterior links having one mean and interior links another. Surprisingly, the statistics of natural link lengths, or even of stream lengths, has not received much investigation beyond the work of Schumm (1956, p. 607-608) on the badlands at Perth Amboy.

The total length L of channels in a basin of magnitude μ is

$$L = l(2\mu - 1), \quad (10a)$$

where l is the link length, which is assumed constant. Similarly, the area A of the basin may be written

$$A = \kappa l^2 (2\mu - 1), \quad (10b)$$

in which the dimensionless coefficient κ is a constant if, as will be assumed, the drainage density

$$D = L/A = 1/(\kappa l) \quad (10c)$$

$\frac{1}{2}L$

is uniform, as would be the case in mature topography developed in a homogeneous environment. Elimination of l between (10a) and (10c) gives κ in terms of commonly measured network parameters,

$$\kappa = (2\mu - 1)/LD. \quad (10d)$$

For the 11 networks tabulated by Melton which are definitely free of geologic controls (1957, Table 2, facing p. 88; Chinle Badlands N and S, Finley and Adams Canyon I and II, Sycamore Canyon I and II, Whipple Basin, Mesa Gulch, Dory Hill Basin, Peters Dome Basin, and Cerro Pavo Basin), the mean value of κ is 0.96. Thus, from (10b) the average basin area drained by each link is approximately that of a square of side l ; and from (10c) the constant of channel maintenance (Schumm, 1956, p. 607) is approximately equal to the mean link length. Assuming $\kappa = 1$, $\mu \gg 1$, and noting from (9c) that the bifurcation ratio is approximately $\frac{4}{3}$, so that (Shreve, 1966, p. 21)

$$\mu \approx 4^{\omega-1}, \quad (10e)$$

where μ and ω are the basin magnitude and order, respectively, and \approx denotes approximate equality, leads to a "law of basin areas,"

$$A \approx (2/D^2)4^{\omega-1}. \quad (10f)$$

The basin-area ratio is thus $\frac{4}{3}$, in good agreement with observation (e.g., Schumm, 1956, p. 604-605). *agree! ie. $R_A = R_B$*

Because from (9c) the bifurcation ratio is approximately $\frac{4}{3}$, a simple summation for $\mu \gg 1$ shows that the total number S_S of Strahler streams in networks of magnitude μ will on the average be close to $4\mu/3$; hence, for all networks of sufficient size, the average ratio S_S/μ will be approximately $\frac{4}{3}$, in agreement with the mean observed ratio of 1.34 for the 11 networks free of geologic controls tabulated by Melton, whose magnitudes range from 19 to 111.

Once more assuming $\kappa = 1$, $\mu \gg 1$, defining $F_S = S_S/A$, and using (10b) and (10c), leads to the approximate relationship

$$F_S/D^2 \approx \frac{2}{3}. \quad (11a)$$

This may be compared with the empirical equation

$$F_S/D^2 \approx 0.694 \quad (11b)$$

obtained by Melton (1958, p. 36-37) by analysis of data from 156 drainage basins. For the 11 networks previously considered, the mean value of F_S/D^2 is 0.71.

In terms of links rather than Strahler streams, the relationship is exactly

$$F/D^2 = \kappa, \quad (12)$$

where F is the average number of links per unit area. This equation is tautological, inasmuch as it follows directly from the definitions of F , D , and κ ; it does not depend upon constant link lengths, uniform drainage density, or topologically random networks. Similarly, the ratio F_S/D^2 investigated by Melton (1958, p. 37) is exactly equal to κ divided by the number of links per Strahler stream. Part of the scatter in his diagram of F_S versus D is therefore due to variations in network topology, as he recognized (Melton, 1958, p. 37, 38, 43-46); and the remainder is due to fluctuations in κ . Use of the channel link (a theoretical concept unrecognized at the time) rather than the Strahler stream as the basic channel unit would have eliminated the scatter due to topological variations and reduced the problem to investigation of the behavior of the conceptually simple quantity κ .

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hypothesis concerning the lengths of links. The simplest, and one that does not appear to be too far from the truth, is that all of the links have the same length. In this case, the "law of stream orders" corresponding to the previous hypothesis would state that the average length of streams increases with order as a geometric series with ratio 2. In like manner, the total length of streams in basins of order μ would increase with successive orders approximately as a geometric series with ratio 4, in good agreement with observations (e.g., Schumm, 1956, p. 604-605).

As alternative to these geomorphological hypotheses, they are based upon averages over a finite population and so, like Horvath's (Shreve, 1966, p. 17), do not require complete distributions needed for a statistical analysis of observations made on a large population. Moreover, the hypothesis of constant link length amounts to using a constant link length. A better hypothesis, perhaps, might be that the link lengths follow a log-normal distribution (M. A. Melton, unpublished analysis; Schumm 1956, p. 607-608) with mean inversely proportional to stream order, or perhaps with exterior links having one mean and interior links having another. Surprisingly, the statistics of natural stream lengths, or even of stream lengths received much investigation beyond the work of Schumm (1956, p. 607-608) on the networks at Perth Amboy.

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